

# Hours Worked in Europe and the US: New Data, New Answers\*

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## **Abstract**

We use national labor force surveys from 1983 through 2015 to construct hours worked per person on the aggregate level and for different demographic groups for 18 European countries and the US. We apply a harmonization procedure to measure hours worked consistently across countries and over time. In the recent cross-section, Europeans work 14% fewer hours than US Americans. Differences in weeks worked and in the educational composition each account for one quarter to one half of this gap. Lower hours per person than in the US are in addition driven by lower weekly hours worked in Scandinavia and Western Europe, but by lower employment rates in Eastern and Southern Europe.

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## I. Introduction

An active recent literature has documented large differences in the levels and trends of aggregate hours worked per person across OECD countries, and specifically lower aggregate hours worked per person in Europe than in the US. This literature traces these lower hours in Europe back to, amongst others, labor income taxation (e.g., [Prescott, 2004](#); [Rogerson, 2006](#); [Faggio and Nickell, 2007](#); [Olovsson, 2009](#); [McDaniel, 2011](#); [Bick and Fuchs-Schündeln, 2017](#); to name a few), institutions ([Alesina et al., 2005](#); [Faggio and Nickell, 2007](#)), and social security systems ([Erosa et al., 2012](#); [Wallenius, 2013](#); [Alonso-Ortiz, 2014](#)). One basic step to understand the causes of the large differences in labor supply is to analyze, first, whether these differences exist for all margins of labor supply, and, second, in how far they are driven by different characteristics of the population. Are fewer Europeans working, or do they work fewer hours per work week, or do they simply enjoy more vacation days? Are the aggregate differences driven by different sectoral compositions across countries, different age compositions, or different educational compositions? Are all Europeans working less than US Americans or is this the case only in specific sectors, for specific age groups, for specific education groups, or for men or women? These questions have been raised early on in [Rogerson \(2006\)](#) as an important avenue for advancing the research agenda on understanding differences in hours worked across countries, but could only be answered to a limited extent based on existing data readily available to researchers. E.g., while [Alesina et al. \(2005\)](#) and [Faggio and Nickell \(2007\)](#) provide evidence on the role of employment, weekly hours, and vacation days in accounting for differences in hours worked per person across countries, they only do so for the year 2004 – the only year for which the OECD released data on average weekly hours worked and weeks worked. Moreover, based on these data they cannot investigate the role of the demographic composition or heterogeneous behavior by different demographic groups for differences in hours worked per person.<sup>1</sup> This paper aims to fill this gap.

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<sup>1</sup>[Faggio and Nickell \(2007\)](#) analyze employment rates by age and gender, but do not have information on employment rates by education and on heterogeneity in hours worked per employed.

This paper constructs “new data” on hours worked per person for the US and 18 European countries from national labor force surveys for the years 1983 through 2015.<sup>2</sup> While using labor force survey data for cross-country studies of employment and hours worked is not “new”, our contribution is to apply the harmonization procedure suggested by [Pilat \(2003\)](#), and to systematically analyze its importance. This harmonization procedure addresses the two main challenges for measuring hours worked consistently across countries and over time from labor force surveys. First, a key difference between the several surveys, at a given point in time across countries or for a given country over time, regards the sampling of reference weeks over the year. The reference weeks range from one specific week only, to a quarter, to the sampling of all 52 weeks of the year. We document that this is not a major problem for the consistent measurement of employment rates. But without any further adjustment hours worked per employed are not measured consistently, because they exhibit a fair amount of seasonality, mostly driven by the uneven distribution of vacation days over the year. Second, we show that even if all 52 weeks of the year are covered, vacation days are severely underestimated in labor force surveys, partly because not all weeks are sampled with equal probability. E.g., the Christmas week is typically sampled much less frequently than other weeks. Therefore, even exploiting continuous year surveys and computing annual hours worked as the average of actual hours over all 52 weeks of the year leads to a significant upward biased estimate. Moreover, the size of the bias exhibits a substantial cross-country variation. We implement a measurement strategy of hours worked per employed that is able to deal with both of the above issues by adjusting for vacation days, i.e. annual leave and public holidays, obtained from external data sources.

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<sup>2</sup>These data are available for download from our webpages. [Ohanian and Raffo \(2012\)](#) also add new evidence on hours worked for OECD countries by constructing a time-series of aggregate hours at the quarterly frequency. [Bick et al. \(2018\)](#) measure hours worked for 80 countries in the early 2000s focusing on the differences between low-, middle- and high-income countries rather than the differences among the high-income countries as we do here. [Burda et al. \(2008, 2013\)](#), [Fang and McDaniel \(2017\)](#), and [Bridgman et al. \(2017\)](#) construct cross-country hours worked measures using time-use surveys. The advantage of time use surveys is the precise measurement of time spent over an entire work day or weekend day. The disadvantage – particularly in comparison to labor force surveys – is the much lower frequency and smaller sample size.

We use the harmonized data to document some basic facts for recent aggregate hours worked differences between Europe and the US for the years 2013 to 2015. Hours worked per person are on average 14 percent lower in Europe than in the US, ranging from 7 percent lower hours in Eastern Europe to 25 percent lower hours in Southern Europe. While weeks worked are uniformly lower in Europe than in the US, in Southern and Eastern Europe lower hours are additionally driven by lower employment rates, and in Western Europe and Scandinavia by fewer hours worked per work week. The other component of labor supply always points in the opposite direction: individuals in Southern and Eastern Europe tend to work longer weekly hours than US citizens, and individuals in Scandinavia and to some extent Western Europe are more likely to be employed than US citizens. Thus, there exists a strong negative cross-country correlation between weekly hours and employment rates on the aggregate level among European countries.

The richness of our micro data allows us to move beyond the aggregate hours worked differences between Europe and the US and to provide “new answers” regarding the origins of these differences. Specifically, we investigate how important differences in the demographic composition are for the aggregate patterns, and which groups are most important for driving the aggregate patterns. For instance, if older people worked on average fewer hours than younger people in all countries, and European countries had an on average older population, this could partly account for the lower hours in Europe. We consider heterogeneity among four dimensions: gender, age, education, and sector of employment. This is similar to the exercise by [Blundell et al. \(2011, 2013\)](#), who analyze changes over time in France, the UK, and the US, rather than the cross-section of countries.

With regard to differences in the composition, we find that only the educational composition is an important driver of aggregate differences. While weekly hours worked are similar for the low-, medium- and high-educated individuals, employment rates increase substantially by education in all countries. Since Europe has a higher share of low- and medium-educated individuals than the US, especially in Eastern and Southern Europe, the educational composition alone accounts for

one quarter to one half of the Europe-US hours differences, about as much as differences in weeks worked do. However, even conditional on controlling for the demographic composition the basic patterns concerning employment rates and weekly hours worked prevail: lower hours per person in Scandinavia and Western Europe than in the US are in addition driven by lower weekly hours worked, whereas lower hours per person in Eastern and Southern Europe are in addition driven by lower employment rates.

Are these remaining differences in employment rates and weekly hours worked caused by specific demographic subgroups, or are they prevalent for all groups in a country? We find that for Scandinavia, Western Europe, and Southern Europe the basic patterns are quite homogeneous across groups: in Scandinavia and Western Europe, all groups tend to exhibit high employment rates and low weekly hours compared to their US counterparts. However, being female increases the differences to the US in both margins. In Southern Europe, all demographic groups tend to exhibit both lower employment rates and lower weekly hours than their US counterparts, with the exception of the low educated. Eastern Europe features more heterogeneity across the groups than the other regions.

The remainder of the paper is structured as follows: Section II describes the micro data sets and explains how we harmonize the data across countries and over time. It also explains how we measure hours worked per person, employment rates, and hours worked per employed, and compares them to the corresponding figures from the National Income and Product Accounts. Section III presents the aggregate hours worked differences between Europe and the US in our data, and documents the cross-country relationships between the three components of labor supply, i.e. the employment rate, weeks worked, and weekly hours worked. In Section IV we conduct our main analysis and investigate the role of different demographic subgroups in accounting for the Europe-US hours difference. Finally, Section V concludes with potential implications of our findings for modeling and suggestions for future work.

## II. Data and Measurement

In this section, we first introduce the country-specific labor force surveys, followed by the key questions from these surveys that we use to measure hours worked per person. Sections II.3 and II.4 then present the two main challenges that we face in measuring hours worked consistently across countries and over time, namely the sampling of reference weeks and the underreporting of vacation days. Only the former affects the measurement of the employment rate (by definition), but we show that this is not a major problem. In contrast, we document that both of these challenges play a potentially large role for the measurement of hours worked per employed. Section II.5 explains how we construct hours worked per employed to overcome these two issues.

At the start of this section, let us define the main variables used throughout the paper.  $H$  refers to average annual hours worked per person, and  $H^E$  are average annual hours worked per employed.  $e$  is the employment rate,  $h$  are hours worked per employed in a non-vacation week (a term which we use interchangeably with weekly hours worked), and  $w$  is the average number of non-vacation weeks in a year (which we use interchangeably with weeks worked), such that

$$H^E = h \times w,$$

and

$$H = e \times H^E = e \times h \times w.$$

We call the employment rate  $e$  and hours worked per employed  $H^E$  the two margins of labor supply, as usual in the literature, while we talk about components when we further divide hours worked per employed into weekly hours in a non-vacation week and weeks worked.

## *Data Sets*

We construct hours worked from two different micro data sets, namely the European Union Labor Force Survey and the Current Population Survey.

*European Union Labor Force Survey* The European Union Labor Force Survey (EU LFS) is a collection of annual labor force surveys from different European countries. We use the yearly surveys, since the quarterly ones do not provide information on education. The EU LFS covers Belgium, Denmark, France, Germany, Greece, Italy, Ireland, the Netherlands, and the UK from 1983 on, Portugal and Spain starting in 1986, Austria, Norway, and Sweden starting in 1995, Hungary and Switzerland starting in 1996, and the Czech Republic and Poland starting in 1997.<sup>3</sup> The sample size of the EU LFS varies across countries and also within a country over time, but is always of considerable magnitude. The minimum annual sample size in the original data we use is 15,400 for Denmark, a country with roughly 5.5 million inhabitants.

*Current Population Survey* For the US, we use the Current Population Survey (CPS), which is a monthly survey of around 60,000 households. Specifically, we work with the CPS Merged Outgoing Rotation Groups data provided by the National Bureau of Economic Research (see <http://www.nber.org/data/morg.html>). This data set includes only those interviews in which the households are asked about actual and usual hours worked, namely the fourth and eighth interview of each household. The data cover around 300,000 individuals per year.<sup>4</sup>

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<sup>3</sup>For the Netherlands, we have information from 1983, 1985, and annually from 1987 on. Prior to 1991 the data for Germany cover only West Germany. The EU LFS also covers Finland from 1995 on. However, the Finnish data have large numbers of missing observations for several years, which implies that we could only use data from 1997 to 2002 for our analysis. The EU LFS covers more transition countries, which we exclude from the analysis because of a shorter time series.

<sup>4</sup>While it is well known that wages are imputed for many individuals in the CPS, this is not the case for hours, of which less than 1% are imputed.

### ***Key Survey Questions and Sample Selection***

To measure employment rates and hours worked per employed person, we rely on five survey questions. We construct the employment rate based on the self-reported employment status in the reference week, which is usually the week prior to the interview. Employed individuals include, in addition to employees, the self-employed and unpaid family workers. Moreover, being employed does not require that the individual works positive hours in the reference week. To estimate hours worked for employed individuals, we rely on questions about actual hours worked in the main job in the reference week, actual hours worked in additional jobs in the reference week, usual hours worked in the main job in a working week, and reasons for having worked more or less hours than usual in the reference week. In principal, the surveys should cover all hours worked in both the formal and informal sector. Online Appendix Section A.1.1 discusses how we deal with missing values for these questions, explains why we have to drop three country-year pairs (1983 for Denmark, 2001 for the UK, and 2005 for Spain), and provides an overview of the annual final sample size per country, which ranges from 10,000 to 450,000 observations, with an average of 115,000 observations. Our final sample includes all observations for individuals aged 15 to 64.

### ***Challenge 1: Sampling of Reference Weeks***

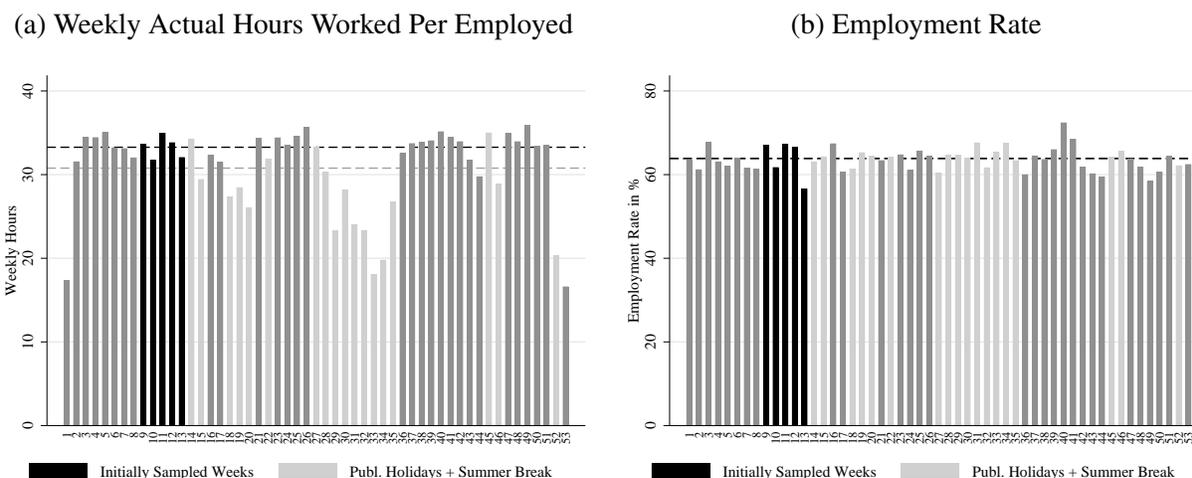
We face two key challenges when measuring hours worked per person across countries: first, the sampling of reference weeks differs across countries and within European countries over time, potentially inhibiting the comparability of measured hours. Secondly, we find that vacation days are systematically underreported in labor force surveys.

The main difference between the several surveys, at a given point in time across all countries or for a given country over time, regards the sampling of reference weeks over the year. The CPS samples all 12 months of the year, but uses as a reference week always the week including the 12th day of a month. Therefore, most major US public holidays are not captured by the CPS (e.g.

Memorial Day, 4th of July, Thanksgiving, Christmas Day). The reference weeks in the national labor force surveys of the European countries initially fell only into country-specific periods ranging from one single reference week to the sampling of half a year. From 2005 onwards, each week of a year is sampled in the European Labor Force Survey for most countries. Eurostat, in its efforts to harmonize the different surveys as much as possible, treated the changes in reference weeks in a two-step procedure. First, if the change to sampling all weeks occurred in a country before 2005, the EU LFS micro data reflect this by changing from sampling only single weeks to sampling the second quarter of the calendar year (April to June) from then on until 2004. Exceptions to this rule are detailed in Online Appendix A.1.2. In a second step from 2005 onwards, when the majority of countries included in the EU LFS had switched to continuous surveying, Eurostat makes all 52 weeks of the year available. The only exceptions to this second step rule are the UK (continuous surveying from 2008 on), Ireland (from 2009 on), and Switzerland (from 2010 on).

The differences in the sampling of reference weeks raise the question in how far estimates of the employment rate and hours worked per employed are comparable across countries in a given year, and within a country over time. Since there was no change for the US, the latter question is only relevant for the European countries. To give a concrete example, Figure 1 shows weekly average actual hours worked per employed in the left panel, and weekly employment rates in the right panel, for France in 2015 by reference week. The black weeks are the weeks that were initially sampled in France until 2002 (weeks 9 to 14). The light gray weeks are weeks that include a public holiday, plus the summer vacation weeks. As one can see in the left panel of Figure 1, weekly hours worked vary substantially over the 52 weeks of the year, and are on average significantly lower in vacation weeks. The sampling of reference weeks thus matters for the aggregate: the gray dashed horizontal line indicates average weekly hours worked over all 52 weeks, while the black dashed horizontal line indicates average weekly hours worked over the initially sampled black weeks: the difference amounts to 2.7 weekly hours (8.6 percent). By contrast, the right panel of Figure 1 shows that the employment rate varies much less over the year. The average annual employment

Figure 1: Avg. Weekly Actual Hours Worked Per Employed and Employment Rate (France, 2015)



Note: Each bar shows the average (a) weekly hours worked per employed and (b) employment rate for individuals aged 15-64 for each week of the year in France for the year 2015. The black dashed line shows the annual average using only the initially sampled weeks (black bars), the gray dashed line shows the annual average calculated using all weeks of the year.

rate is virtually independent of whether only the initially sampled weeks or all 52 weeks are used (the gray and black dashed horizontal lines overlap).

To analyze the extent of seasonality in measures of the employment rate and hours worked per employed more generally, we exploit (as in the example above for France) the fact that the EU LFS from 2005 onwards samples all weeks of the year for most countries, and also states the reference week for each interviewed person. Thus, for these years, we analyze how employment rates and hours worked per employed differ if they are constructed using information from only the initially sampled weeks, i.e. only those weeks which served as country-specific reference weeks prior to the change to continuous surveying (e.g. the black weeks in Figure 1 for France), and using all weeks of the year as reference weeks. We use all available country-year pairs from 2005 to 2015, which are in total 186 observations.<sup>5</sup>

<sup>5</sup>For 15 of our 18 European countries we can do the comparison in all years starting 2005, for the UK, Ireland, and Switzerland only starting later. In Online Appendix A.1.2, we show corresponding evidence when the second quarter is used as reference period, as done for the interim period in the European countries. Differences to using all weeks of the year as reference period are somewhat smaller than under the initially sampled weeks, but still substantial.

Table 1: Absolute %-Deviations: Initially Sampled Weeks vs. All 52 Weeks of the Year

	Mean	75th	90th	95th	99th	Max
$H^{E,raw}$	3.6	5.0	7.4	8.0	11.4	12.6
$H^E$	0.8	1.1	1.5	2.1	4.1	5.0
$e$	0.6	0.7	1.1	1.5	5.3	5.9
$H^{raw}$	3.4	4.7	7.7	8.5	12.0	12.6
$H$	1.4	1.8	2.7	3.1	7.1	8.6

Note: For  $H^{E,raw}$ ,  $H^{raw}$ ,  $H^E$  and  $H$  we report deviations in percent, whereas for  $e$  we use percentage points. All three measures refer to the population 15-64 and are calculated for all available European country-year pairs from 2005 to 2015 (in total 186 observations).

*Hours Worked per Employed* We calculate a measure of “raw” annual hours worked per employed as the sum of actual weekly hours worked of all employed individuals multiplied by 52 weeks in a year:

$$H^{E,raw} = 52 \times \frac{1}{N^e} \sum_{i=1}^N \varepsilon_i \eta_i^{raw},$$

where  $\varepsilon_i$  is the self-reported employment status of individual  $i$ , which takes the value 1 for anyone reporting to be employed (including self-employment or being an unpaid family employee), and 0 otherwise;  $\eta_i^{raw}$  are the actual hours worked in the reference week in all jobs of individual  $i$ ;  $N$  is the total number of individuals; and  $N^e$  is the number of employed individuals, i.e.  $N^e = \sum_{i=1}^N \varepsilon_i$ . All measures are calculated using the individual survey weights, which we omit here for the ease of exposition.

Table 1 shows in the first row the absolute percent deviation of annual hours worked per employed when using only the initially sampled reference weeks from using all weeks of a year, which amounts on average to 3.6 percent. At the 90th percentile, the deviation already reaches 7.4 percent, with a maximum deviation of 12.6 percent in Sweden for the year 2008. Thus, the sampling of the reference weeks matters substantially for the measurement of hours worked per employed. With only few exceptions, the hours when only specific survey weeks are sampled are

larger than when all weeks of the year are sampled.<sup>6</sup>

*Employment Rate* The employment rate ( $e$ ) is simply given by the number of employed individuals divided by all individuals in the sample

$$e = \frac{N^e}{N}. \quad (1)$$

Equation (1) is the employment to population ratio for the population 15-64, but we henceforth refer to it as the employment rate.<sup>7</sup> The third row of Table 1 shows that the employment rate exhibits significantly less seasonality than the raw hours worked per employed measure. The average absolute percentage point deviation between constructing the employment rate only based on specific survey weeks relative to the entire year amounts to only 0.6 percentage points. At the 90th percentiles, the difference is still relatively small with 1.1 percentage points. Moreover, employment rates are not systematically higher (or lower) if only specific weeks are sampled than if the entire year is sampled. There are only two country-year pairs for which the deviations are larger than 5 percentage points, namely from Germany and Belgium. Both countries used only one specific week as reference week (Germany until 2004, Belgium until 1998), suggesting that the corresponding time-series of the employment rates have to be used with some caution.

*Hours Worked per Person* The fourth row of Table 1 finally shows the seasonality of hours worked per person ( $H^{raw}$ ), which is the product of hours worked per employed and the employ-

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<sup>6</sup>This is the case for all countries in all but at most two years. The only exception are the Netherlands, where hours are lower in the initially sampled weeks in seven out of eleven years.

<sup>7</sup>Online Appendix Section A.1.3 reports two alternative measures of the employment rate. The first alternative measure relies on defining employment based on usually working positive hours. This leaves the employment rate virtually unchanged (see Table A.6). The second alternative measure involves a different definition of employment for women on maternity leave. This has only a modest impact on female employment rates, and is far too small to drive international differences in female employment rates (see Table A.7). Note that by construction, hours worked per person remain always the same no matter what definition is used for the employment rate, because an increase (decrease) in the alternative employment rate relative to our baseline definition is always offset by a decrease (increase) in the corresponding measure of hours worked per employed.

ment rate. The mean absolute deviation in hours worked per person between the initially sampled weeks and all weeks of the year amounts to 3.4 percent, and at the 90th percentile it already reaches 8.5 percent.<sup>8</sup> Thus, the strong seasonality of hours worked per employed is fully reflected in hours worked per person.

### ***Challenge 2: Underreporting of Vacation Weeks***

For most European countries all weeks of the year are sampled from 2005 onwards. If each week would be sampled with equal probability and cover a representative sample of the population week by week (rather than only over the entire year), then average vacation weeks should be captured accurately in the labor force surveys, and one could compute average annual hours worked simply by summing up average weekly hours worked in each week over all 52 weeks of the year. Calculating the average number of vacation weeks from the micro data yields an average of 3.4 “self-reported” vacation weeks in Europe, defining vacation as the sum of public holidays and annual leave.<sup>9</sup> Given that public holidays alone in all countries sum up to 1.5 to 2.5 weeks, these self-reported weeks of annual leave and public holidays seem implausibly low. In fact, based on external data sources (national statistical offices, unions, employer organizations, the European Industrial Relations Observatory, and the International Labor Organization), the average weeks of vacation over the same time period for our sample of European countries amount to 6.8 weeks per year.<sup>10</sup> Put differently, even exploiting continuous year surveys and computing average actual hours over all 52 weeks of the year without any further adjustment ( $H^{E,raw}$ ) leads to a significant overestimate of annual hours worked. Moreover, the size of the bias exhibits a substantial cross-country variation.

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<sup>8</sup>The correlation between  $H^{E,raw}$  and  $e$  is slightly negative (-0.045), which leads to slightly smaller absolute percent deviations for  $H^{raw}$  than for  $H^{E,raw}$ , except at the extremes.

<sup>9</sup>Table A.1.1 in Online Appendix A.2.2 reports the results for each country as well as the description of how we construct the self-reported vacation weeks.

<sup>10</sup>Online Appendix A.2.1 lists the exact external data sources and explains further assumptions in the construction of the number of vacation weeks. It also shows that the time series variation in vacation weeks within each country is small, with the notable exception of Denmark. For the US, the discrepancy between self-reported and external vacation weeks is somewhat smaller but still substantial, with 1.3 vs. 3.7 weeks. We expect a discrepancy for the US, since weeks with major public holidays are not sampled.

To get a better understanding of the sources of this discrepancy, we investigate in more detail the case of Germany, which features the largest difference among all countries, amounting to 5.1 weeks. While we provide all details in Online Appendix Section A.2.2, our results of this analysis can be summarized as follows. First, using the German Socio-Economic Panel, [Schnitzlein \(2011\)](#) reports that on average 3 of the average 31 entitled days of annual leave per year go unused. Thus, the underusage of entitled leave can explain only a very small portion of the discrepancy of 5.1 weeks between self-reports and official vacation days. Second, not all weeks are sampled with the same probability in the EU LFS, see Figure A.1 in Online Appendix Section A.2.2. For example, the reference week which contains Christmas day is on average sampled with a probability of 0.7%, significantly below the 1.9% that would be implied by equal sampling. Third, the German Statistical Office also has some evidence that respondents might dislike using a vacation week as a reference week, either because they are too busy the first week after a vacation to fill out the questionnaire, or because they perceive it as “inappropriate” to use a vacation week as reference week. Summarizing, at least for Germany there exists evidence of underreporting of days of annual leave and public holidays in the EU LFS even after the introduction of continuous sampling over the entire year. It seems at least not implausible that these factors can explain the discrepancies between the self-reported and external vacation weeks for the remaining countries as well. Therefore, it is important to adjust hours worked per employed for vacation using external data even when analyzing the recent cross-section from 2005 onwards, when all weeks of a year are sampled.

### ***Measurement of Hours Worked per Employed***

As documented so far, the raw measure of hours worked per employed suffers from two weaknesses: first, seasonality in hours worked per employed impedes the comparability of hours worked across countries and within European countries over time because of differences in the sampled reference weeks; second, vacation weeks are underreported in the labor force surveys even if all

weeks of a year are sampled. We therefore apply the following adjustment to construct a measure of hours worked per employed that overcomes both weaknesses: we first obtain individual hours worked in a *non-vacation* week, i.e. a typical work week without a reduction in work time because of a public holiday or annual leave, and then require *consistency* of annual leave and public holidays in the micro data with the country-wide average. This closely follows [Pilat \(2003\)](#).

To calculate hours worked in a *non-vacation* week  $\eta_i$  for each employed individual, we use as a baseline actual hours worked in all jobs in the reference week. However, if a respondent indicates that he/she worked less hours than usual in the main job in the reference week, and states as the main reason for doing so public holidays or annual leave, we replace weekly hours by usual hours in the main job plus actual hours worked in all additional jobs.<sup>11</sup> Thus,

$$\eta_i = \begin{cases} \text{usual hours in main job} & \text{if actual hours in main job} < \text{usual hours in main job} \\ + \text{ actual hours in all additional jobs} & \text{because of annual leave or public holiday} \\ \text{actual hours in all jobs } \eta_i^{raw} & \text{otherwise.} \end{cases}$$

We refer to these hours  $\eta_i$  as “hours worked in a non-vacation week”. Averaging over our population of interest yields mean weekly hours worked in a non-vacation week  $h$ , i.e.

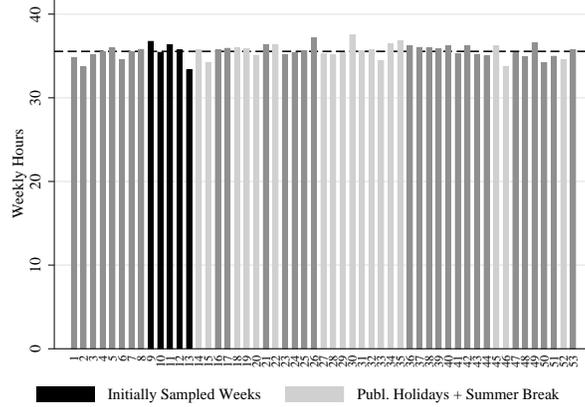
$$h = \frac{1}{N^e} \sum_{i=1}^N \varepsilon_i \eta_i. \quad (2)$$

If people work less hours than usual for other reasons than vacations, e.g. because of sickness, or more hours than usual because of overtime, this is captured by this measure. [Figure 2](#) shows

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<sup>11</sup>Note that respondents can only indicate one reason for working different hours than usual, and for additional jobs we do not have information on usual hours. One might be worried that individuals with multiple jobs might take vacation from their main job to work more hours in their additional job. However, in the US only about 6% of the employed population hold multiple jobs, and about 3% on average in the European countries in our sample. Hence, any potential effects on our measure of aggregate hours worked should be negligible. In [Online Appendix A.1.4](#) we discuss some differences between the CPS and EU LFS questionnaire regarding the construction of our hours measure, which however have virtually no impact on the statistics presented in the paper.

Figure 2: Average Hours Worked in a Non-Vacation Week  $h$  (France, 2015)



Note: Each bar shows average hours worked in a non-vacation week for employed individuals aged 15-64 years for each week of the year in France for the year 2015. The black dashed line shows the annual average using only the initially sampled weeks (black bars), the gray dashed line shows the annual average calculated using all weeks of the year.

this measure of hours worked in a non-vacation week for France in 2015. It does hardly show any seasonality anymore, in contrast to the raw measure of average weekly hours worked in the left panel of Figure 1. As a consequence, the average annual hours worked per employed do not depend on whether only the initially sampled weeks or all 52 weeks are used (the gray and black dashed horizontal lines overlap).

In order to establish consistency of annual leave and public holidays in the micro data with the country-wide average from external data sources, we multiply mean weekly hours worked in a non-vacation week  $h$  with the number of non-vacation weeks  $w$ . We obtain our estimate of  $w$  by subtracting the country-wide average weeks of annual leave and public holidays reported in external data sources (for details see Online Appendix A.2.1 from the 52 weeks of a year, i.e.

$$w = 52 - \frac{\text{country-wide average days of annual leave \& public holidays}}{5}. \quad (3)$$

The external data sources provide vacation days, which we convert into vacation weeks by dividing by 5, assuming 5 work days per week. Note that this adjustment works equally well for part-time

and full-time workers. Part-time workers get the same number of vacation *weeks* as full-time workers; the number of vacation *hours* are adjusted according to their part-time schedule. The disadvantage of this procedure is that we rule out any heterogeneity in the number of days of annual leave and public holidays. One implication of this is that individuals who work while being on vacation, and thus only use a fraction of their vacation time, are assumed to take the full vacation time as given by the country-wide average. Since we only observe each individual for one week of the year, we cannot verify whether this assumption holds. To be precise, for individuals reporting positive hours worked but less than usual hours because of vacation, we cannot distinguish whether they took some days off in the reference week and only worked on the remaining days, or whether they took the full week off but worked some hours while on vacation. To the extent that the latter is the case, we are underestimating hours worked per person. In Online Appendix Section A.2.3, we conduct a robustness check which allows for heterogeneous number of vacation weeks by demographic subgroups. Specifically, we assume that the heterogeneity in self-reported vacation weeks reflects the true degree of heterogeneity, but we still impose that the average number of vacation weeks corresponds to the one from external sources. Allowing heterogeneity in vacation weeks has only a negligible effect on our results.

The second row of Table 1 shows the deviation of the annual hours worked per employed measure ( $H^E = h \times w$ ) if only the initially sampled survey weeks are used as reference weeks, compared to all weeks of the year. While for the “raw” measure of hours worked per employed in the first row the mean absolute percent deviation amounts to 3.6 percent, the mean absolute percent deviation of our measure of hours worked per employed is with 0.8 percent substantially smaller. Thus, externally adjusting for vacation is successful in insuring comparability of the data over time. This becomes apparent for the example of France in the year 2015 by comparing Figures 1a and 2. Only at the top end of the distribution, we still see a non-negligible effect of the sampling of the reference weeks. The three countries with deviations in the top five percent are Belgium, Norway, and Poland. As the last row of Table 1 shows, the deviations in hours worked per person

between the initially sampled weeks and all weeks of the year is also significantly reduced by our adjustment procedure. The mean absolute deviation amounts to 1.4 percent, compared to 3.4 percent for the “raw” measure of hours worked per person. At the 95th percentile, the absolute deviation is 3.1 percent, compared to 8.5 percent without the adjustment.<sup>12</sup>

Finally, while we document the need to adjust for hours lost due to public holidays and annual leave, the same could in principle apply to working less hours than usual because of sick leave or other reasons, and also to working more hours than usual (overtime and additional jobs). There are two dimensions to this issue. First, in how far do such differences between usual and actual hours vary with the set of reference weeks available? In Online Appendix Section A.3, we provide evidence that none of the other reasons for working more or less time than usual vary with the set of reference weeks. Put differently, with regard to these categories the initially sampled weeks seem to be representative for the average over all weeks. Second, is there systematic under- or overreporting of these categories? Unfortunately, sick days from external data sources are only available for a subset of countries and years. For those country-year pairs, we find that the self-reported number of sick days in our data are smaller than sick days from external data sources. This is similar to our finding for vacation days. However, the discrepancy is on average smaller than for vacation days and public holidays, with on average 1.1 weeks in Europe and 0.3 weeks in the US, see Online Appendix Section A.3.

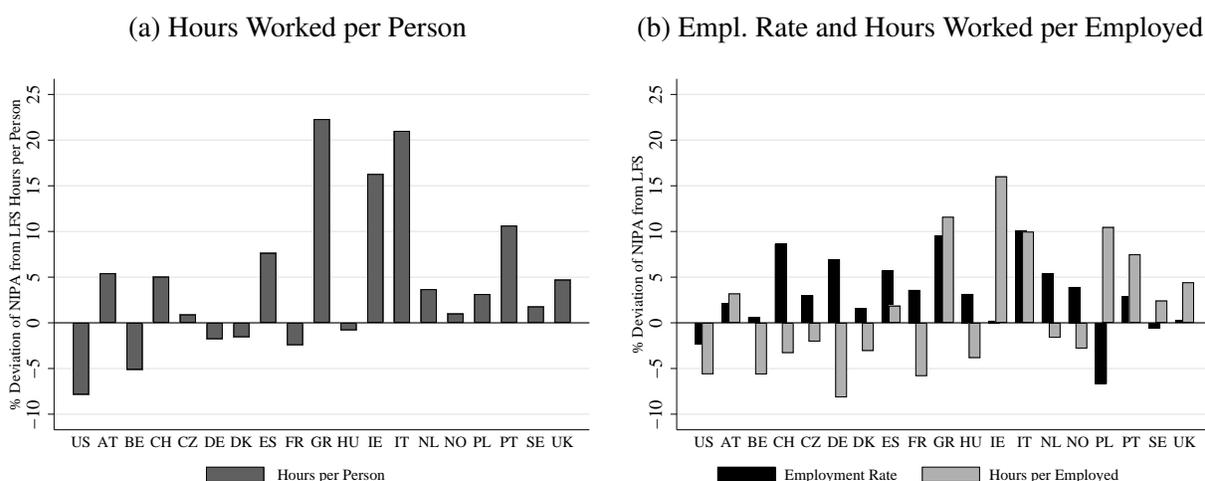
### ***Comparison with NIPA Hours***

As is common practice when working with micro data sets, see e.g. the 2010 Review of Economic Dynamics special issue on “Cross Sectional Facts for Macroeconomists” (Krueger et al., 2010), we compare our aggregate statistics on hours worked per person based on the labor force surveys (LFS) with those from the National Income and Product Accounts (NIPA) as provided by the

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<sup>12</sup>The correlation between  $H^E$  and  $e$  is positive (0.095), which leads to larger absolute percent deviations for  $H$  than for  $H^E$ .

Figure 3: Percent Deviations of Hours Worked and Employment in NIPA from LFS



Note: NIPA hours worked per person are calculated by dividing NIPA total hours from the OECD National Accounts by the population aged 15 to 64 from the OECD ALFS database. For the LFS data, total LFS hours (everyone aged older than 15) are divided by the LFS population aged 15 to 64. The NIPA employment rate is calculated as total employment from the OECD National Accounts (domestic concept) divided by the population aged 15 to 64, while the LFS employment rate is total number of employed (older than 15) divided by the LFS population 15 to 64. All numbers are for the year 2012.

OECD’s “National Accounts Database”.<sup>13</sup> Specifically, we calculate a NIPA measure of hours worked per person by dividing NIPA total hours by the population aged 15 to 64 obtained from the OECD’s “Annual Labour Force Statistics”. For the LFS data, we divide total LFS hours (rather than only those by the population aged 15 to 64 as in the previous subsections) by the population aged 15 to 64 directly taken from the LFS.

Figure 3a shows the average country-specific percent deviation of NIPA estimates for hours worked per person from LFS estimates for 2012, the last year in which all data are available for all countries. On average across all countries, hours worked per person are 43 hours (4.4 percent) higher in the NIPA data than in the LFS data. In 6 out of 19 countries, the NIPA estimates are below the LFS estimates (on average by -3.3 percent).<sup>14</sup> Figure 3b shows that these differences

<sup>13</sup>An alternative data source for aggregate hours worked is the Conference Board’s “Total Economy Database” (TED). For the majority of countries and years in our sample the data from the OECD and the TED are exactly the same, see Online Appendix C.2.

<sup>14</sup>To put these deviations between the NIPA hours and the LFS hours measures into perspective, the implied Europe-US hours worked per person difference (a key statistic our subsequent empirical analysis focuses on) for the year 2012 amounts to -4.5% in NIPA data and -15.9% in LFS data. Hence, our data predict a much larger gap in absolute value

stem from disagreement between the two data sources for both the employment rate and hours worked per employed. In contrast to the cross-section, we show in Online Appendix C.2 that the time trends are generally much more similar between the two different data sources.

What can explain these cross-sectional differences between NIPA and LFS estimates? LFS and NIPA data differ conceptually along two dimensions. First, LFS data cover only civilian, non-institutionalized residents aged 15 and older, while NIPA data do not impose these restrictions to ensure that the labor inputs are consistent with the measurement of gross domestic output (GDP). Second, the NIPA estimates are usually constructed in country-specific ways from multiple data sources (administrative data, social security data, employer surveys, labor force surveys, census data, etc.). Online Appendix A.4 provides more details on both of these conceptual differences and documents that the employment rate differences in Figure 3b cannot be reconciled with differences in the covered population. For hours worked per employed, we cannot conduct such a comparison. For the US, [Abraham et al. \(2013\)](#) document which features of the underlying data sources drive the differences between NIPA and LFS employment rate estimates, and [Eldridge et al. \(2004\)](#) and [Frazis and Stewart \(2010\)](#) do the same for hours works per employed. While the details are specific to the US, these papers highlight that the NIPA estimates are not necessarily superior to the LFS estimates, see Online Appendix A.4 for details. E.g., [Ramey \(2012\)](#) suggests that the LFS hours yield a more accurate picture of the true work hours of individuals.

The combination of multiple data sources might deliver more accurate estimates of employment and hours worked for a given country. The downside is that the cross-country comparability suffers, despite the efforts to harmonize measurement through the Systems of National Accounts, see [Fleck \(2009\)](#). In fact, the OECD remarks on its website that “The [hours worked] data are intended for comparisons of trends over time; they are unsuitable for comparisons of the level of average

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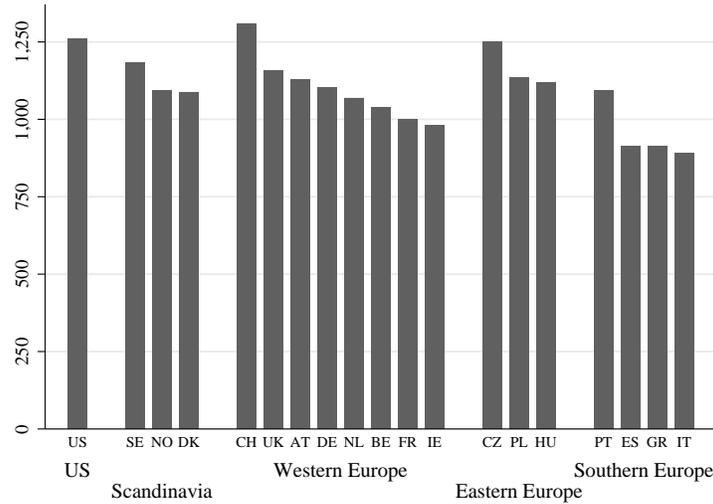
than the NIPA data. Most of the macro literature on hours worked cited in the introduction base their analysis on NIPA data (or some variant of it). Comparing the facts reported in these papers among each other and to the facts reported here is however difficult, since the NIPA hours underwent massive revisions over time. In a companion paper, [Bick et al. \(2017\)](#), we document these revisions and show that they affect the measured Europe-US hours gap tremendously.

annual hours of work for a given year, because of differences in their sources,” and recommends using employment rates based on labor force surveys for cross-country comparisons: “National Labor Force Surveys are the best way to capture unemployment and employment according to the ILO guidelines that define the criteria for a person to be considered as unemployed or employed... While data from LFS make international comparisons easier compared to a mixture of survey and registration data, there are some differences across countries in coverage, survey timing, etc, that may affect international comparisons of labour market outcomes”.<sup>15</sup> Our approach deals with one of the main differences in the cross-country comparability of the LFS, namely the survey timing.<sup>16</sup> In any case, only the LFS give us the opportunity to investigate Europe-US hours differences in detail, since they allow to construct average hours for demographic subgroups.

### **III. The Aggregate Europe-US Hours Difference**

In this section, we start using our harmonized data to document some basic facts for aggregate hours worked differences between Europe and the US for the most recently available years (2013-2015). These facts are hardly influenced by any persistent effects of the global financial crisis and the European debt crisis. Focusing on the pre-crisis years 2005-2007 yields very similar results. We restrict our attention to the population aged 15 to 64. Subsequently, we document the cross-country differences for the three components of labor supply, namely the employment rate, weeks worked, and weekly hours worked.

Figure 4: Average Hours Worked per Person



Note: Average hours worked per person are calculated for the population aged 15 to 64 over the years 2013 to 2015.

### *Hours Worked per Person*

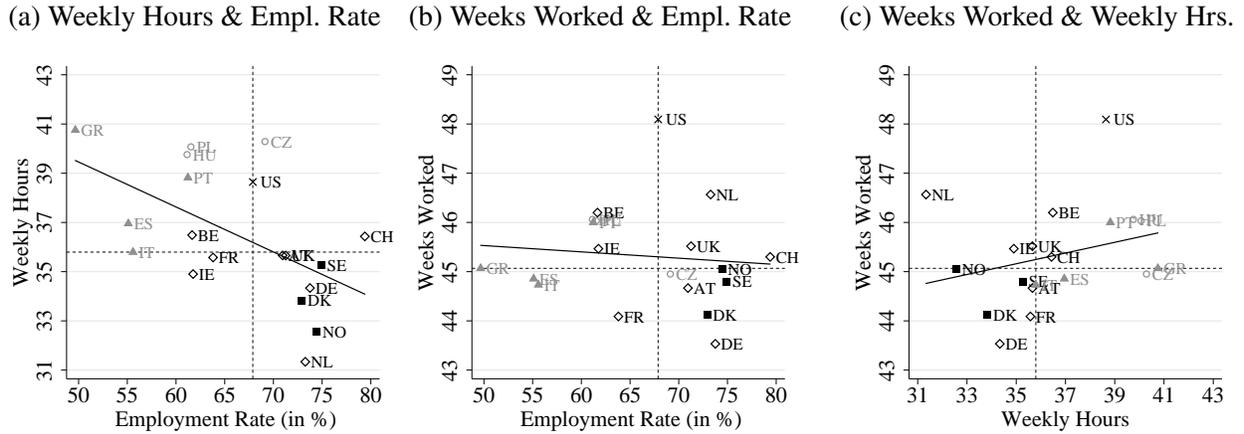
Figure 4 shows hours worked per person in the LFS data on the country level by geographic region. Within each region, we order countries by average hours worked per person. Switzerland, the US, and the Czech Republic stand out with the highest hours worked per person, exceeding 1250, while Italians work less than 900 hours per year. Overall, mean hours worked per person in Southern Europe are the lowest (952 hours), while the mean hours worked per person across the other European regions are quite similar, ranging from 1098 hours in Western Europe to 1169 hours in Eastern Europe. The implied Europe-US hours per person gap is 180 hours, or 14 percent.<sup>17</sup>

<sup>15</sup>Both quotes we retrieved from the OECD's website on August 9, 2017: <http://stats.oecd.org/Index.aspx?DataSetCode=ANHRS> and <http://www.oecd.org/els/emp/basicstatisticalconceptsemploymentunemploymentandactivityinlabourforcesurveys.htm>.

<sup>16</sup>Other reasons impeding the comparability across time and countries of the LFS, which we cannot adjust for, are the revision of population figures used for population adjustment on the basis of new population censuses, as well as changes in the sampling design, and content or order of the questionnaire. For details, see [http://ec.europa.eu/eurostat/statistics-explained/index.php/EU\\_labour\\_force\\_survey](http://ec.europa.eu/eurostat/statistics-explained/index.php/EU_labour_force_survey) (retrieved on August 9, 2017).

<sup>17</sup>We always report simple cross-country averages. Including also adults older than 64 increases the Europe-US hours gap to 19 percent.

Figure 5: Hours Worked Components



Note: The dotted vertical and horizontal lines indicating median values for each variable, and the solid line represents a fitted OLS regression line. All variables are averages over the years 2013 to 2015. Weekly hours worked and the employment rate are calculated for the population aged 15 to 64.

### *The Three Components of Hours Worked per Person*

Average hours worked per person are the product of the employment rate and hours worked per employed, with the latter being the product of weekly hours worked per employed in a non-vacation week and the number of weeks worked per year, i.e.  $H = e \times h \times w$ . Figure 5 shows how these three components relate to each other and reveals interesting patterns of heterogeneity across regions, despite the similar average hours worked per person across European regions. The five regions are marked by different markers and colors, namely the US by a black x, Eastern Europe by a gray circle, Scandinavia by black squares, Western Europe by black diamonds, and Southern Europe by gray triangles.

Figure 5a plots weekly hours worked per employed in a non-vacation week ( $h$ ) against the employment rate ( $e$ ), with the dotted vertical and horizontal lines indicating median values, and the solid line representing a fitted OLS regression line. While hours worked per person do not differ much on average across the European regions, with the exception of Southern Europe as shown in Figure 4, a region-specific pattern for the relationship between the employment rate and weekly hours emerges. Eastern and Southern European countries all have above median weekly

hours with Italy being the median country, but below median employment rates, with the exception of the Czech Republic. The Scandinavian countries in turn have below median weekly hours and above median employment rates. Weekly hours worked per employed in Western Europe are below the median (the exceptions are Belgium and Switzerland), while the employment rates range from below the median in Belgium, Ireland and France to the highest in Switzerland. Taken together, a strong negative cross-country correlation between weekly hours worked per employed and the employment rate of -0.55 arises: countries with high employment rates tend to feature low weekly hours worked per employed in a non-vacation week, and vice versa.

Figure 5b plots the weeks worked ( $w$ ) in a given country against the employment rate ( $e$ ). Weeks worked range in Europe from below 44 in Germany to 46.5 in the Netherlands, with most of the cross-country variation coming from annual leave rather than public holidays (see Online Appendix A.2.1). The US clearly stand out with more than 48 weeks worked per year. Overall, the correlation between weeks worked and the employment rate is weak, amounting to -0.10 for all countries and -0.17 if the US is excluded. Figure 5c plots weeks worked ( $w$ ) against weekly hours worked in a non-vacation week ( $h$ ). The correlation between both variables is 0.28 and drops to 0.21 once the US is excluded, which is consistent with the evidence in [Altonji and Oldham \(2003\)](#). The weak correlations between weeks worked and employment rates or hours worked might suggest that more public holidays or annual leave, i.e. less weeks worked per year, neither induce more people to work nor to work longer hours during a non-vacation week.<sup>18</sup>

#### IV. Decomposing the Europe-US Hours Difference

We now exploit the richness of our micro data and move beyond the aggregate hours worked differences between Europe and the US and provide “new answers” regarding the origins of these differences. Specifically, we are interested in how important differences in the demographic com-

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<sup>18</sup>There exists a strong negative correlation of -0.83 between weekly hours and the part-time rate, defined as the fraction of employed working less than 30 usual weekly hours. Thus, low weekly hours are driven by a large fraction of part-time workers, and vice versa.

position are for the aggregate patterns, and which groups are the main drivers of the aggregate patterns. We proceed in two steps. First, in Subsection IV.1 we conduct a decomposition that analyzes the extent to which different demographic *compositions* and differences in the three components of labor supply contribute to the aggregate Europe-US hours difference. Second, in Subsection IV.2 we analyze whether the stark differences across countries in employment rates and weekly hours worked are driven only by the *behavior* of specific demographic subgroups, or are prevalent across all groups.

### ***Decomposition by Demographics and Components of Hours Worked***

Countries in our sample differ not only in the three components of hours worked, but also in their demographic structure, i.e. the composition of the population by demographic characteristics. We take into account three demographic characteristics, namely gender, age, and education, as well as the sectoral composition of employment in a country. Explicitly accounting for these groups, we write aggregate hours worked per person in country  $c$  as

$$H^c = w^c \times \sum_j^J f_j^c e_j^c \times \sum_k^K h_{j,k}^c s_{j,k}^c, \quad (4)$$

where  $j$  represents a set of demographic characteristics and  $k$  the sector of employment.  $w^c$  are the weeks worked (non-vacation weeks),  $f_j^c$  is the fraction of individuals with a given set of characteristics  $j$  (with  $\sum_{j=1}^J f_j^c = 1$ ),  $e_j^c$  is the employment rate of group  $j$ ,  $h_{j,k}^c$  are the weekly hours worked per employed in a non-vacation week by group  $j$  in sector  $k$ , an  $s_{j,k}^c$  is the share of employed individuals of group  $j$  working in sector  $k$  (with  $\sum_{k=1}^K s_{j,k}^c = 1$ ). Each group  $j$  is defined by the interaction of one of three age groups (15-24, 25-54, 55-64), three education groups (low, medium, high), and gender.<sup>19</sup> Only for the young age group, we allow a fourth educational category “still

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<sup>19</sup>The different levels of education are defined according to the ISCED classifications, with “low” corresponding to lower secondary education, “medium” to upper secondary education, and “high” comprising any tertiary education degree. For example, for the US “low” corresponds to less than high school education, “medium” to completed high school education, and “high” to having at least an associate’s degree.

enrolled in education”. Since only employed individuals can be allocated to a sector, the sectoral composition only affects hours worked per employed, not the employment rate. We consider three broad sectors: services, manufacturing, and agriculture.<sup>20</sup> Last, remember that in the baseline analysis we assume the same number of weeks worked for everyone. In Online Appendix Table B.5 we show that relaxing this assumption has only a negligible effect on our results.

To quantify the contribution of the three components of labor supply, the demographic composition, and the sectoral composition to the difference of the country-specific hours relative to the US, we calculate counterfactual hours by setting for every country one of these five features after the other equal to the corresponding US value. We then measure the incremental fraction of the aggregate hours worked per person difference relative to the US accounted for by this feature.<sup>21</sup> The quantitative results depend on the order in which the features are set country specific. With five features, there are 120 different possible orderings. Here, we report the mean results over all 120 different orderings, and in Online Appendix Table B.4 the minimum and maximum from all 120 different orderings. This decomposition is similar in spirit to the one done by [Blundell et al. \(2011, 2013\)](#) for time-trends in hours in France, the UK, and the US. Note that this is a statistical decomposition, not to be interpreted as a causal analysis, since the different components entering the decomposition are clearly endogenous variables and determined jointly.

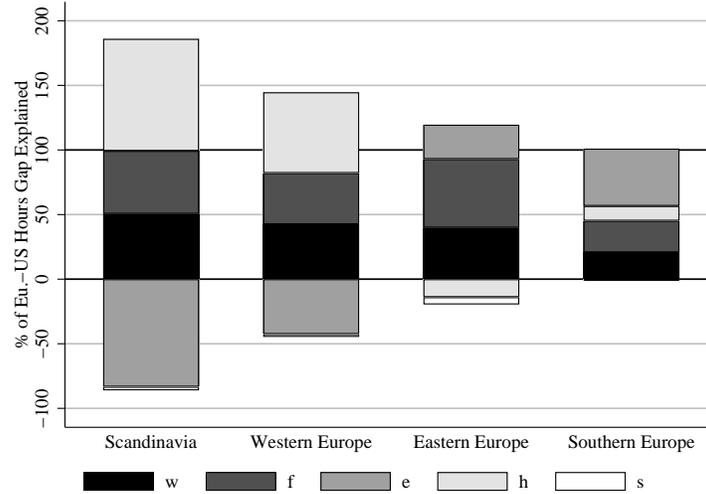
Figure 6 shows the results from our decomposition exercise. Specifically, we show how the difference in hours worked per person to the US is divided into the fraction accounted for by weeks worked  $w$ , the demographic composition  $f$ , employment rates  $e$ , weekly hours  $h$ , and the sectoral composition  $s$ . A negative entry indicates that a factor does not positively contribute to the overall difference, but implies higher hours in the European region than in the US. To facilitate the

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<sup>20</sup>On average across countries and years 2013 to 2015 the educational status is missing for 0.5 percent of individuals and the sector of employment is missing for 0.7 percent of the employed. The corresponding maxima are 3.7 percent for Denmark in 2013, and 7.3 percent for the Netherlands in 2013, respectively. For consistency our decomposition analysis thus features in fact a fifth educational category and fourth sectoral category representing the missing observations.

<sup>21</sup>Online Appendix Section B.1 describes the procedure in detail focusing only on the three components of labor supply analyzed in the previous subsection.

Figure 6: Decomposing the Europe-US Hours per Person Difference



Note: Each part of a bar represents the fraction of the country group-US hours gap for the population aged 15-64 over the years 2013 to 2015 accounted for by the respective component of hours worked. All parts of a bar add up to 100%. The averages for Scandinavia, Western, Eastern, and Southern Europe each exclude the country with the smallest hours worked difference to the US, i.e. Sweden, Switzerland, the Czech Republic, and Portugal, respectively. The exact country-specific values can be found in Online Appendix B.3.

comparison, we show only averages over the regions and show the findings on the country level in Online Appendix B.3.<sup>22</sup>

There are three key findings from the disaggregate decomposition analysis. First, the higher number of vacation weeks in Europe accounts uniformly for between one quarter to one half of the hours difference to the US. Second, differences in the demographic composition constitute another important driving force of hours worked differences between Europe and the US, accounting for around one quarter to one half of the differences, whereas differences in the sectoral composition do not contribute much to the overall difference. Thus, weeks worked and the demographic composition combined could nearly fully account for the hours per person gap with the US in Scandinavia, Western Europe, and Eastern Europe. However, our third finding is that even after controlling for differences in the demographic and sectoral composition, the large differences

<sup>22</sup>We exclude for each region the country with the smallest hours worked difference to the US. For Sweden, Switzerland, and the Czech Republic these small differences in hours worked per person to the US imply extraordinarily large values in the decomposition. Portugal is distinctive from the other Southern European countries because employment rates there indicate substantially higher hours than in the US.

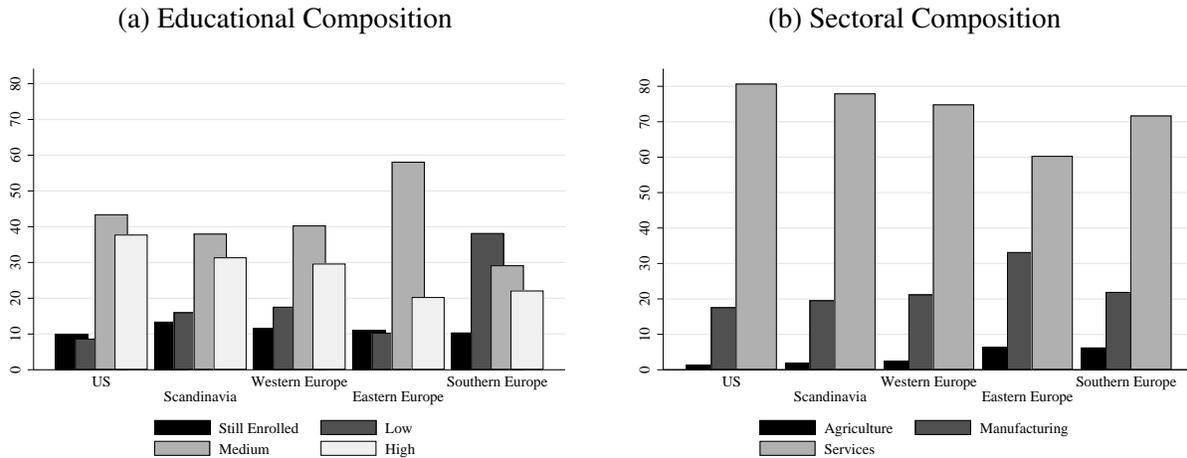
across Europe in the relative contribution of employment rates and weekly hours in accounting for the hours difference to the US persist. For Scandinavia and Western Europe, lower weekly hours worked account for slightly less than 90 percent and slightly more than 60 percent, respectively, of the lower hours per person than in the US. The employment rate differences in turn would predict hours per person differences of a similar magnitude but with the opposite sign. The opposite is true for Eastern Europe, with employment rate differences accounting for roughly one quarter of the hours per person gap with the US, and weekly hours predicting higher hours than in the US. Only in Southern Europe, employment rate (slightly less than 50 percent) and weekly hours (slightly more than 10 percent) differences both contribute positively to the hours gap to the US.

Before investigating which demographic groups differ in their labor supply behavior and are thus driving the remaining differences attributable to employment rates and weekly hours worked, we show why the demographic composition is so important in accounting for the Europe-US hours gap, and why the sectoral composition is not. From a conceptual perspective, the demographic and sectoral compositions matter in the decomposition analysis only if two conditions are met: first, the composition across countries needs to be different; second, the demographic groups need to exhibit either different employment rates or different weekly hours worked *within* a country.

With regard to the composition, since in all sample countries and years women make up around 50 percent of the population aged 15 to 64, applying the gender composition of the US to a country will not affect the corresponding counterfactual estimate of aggregate hours worked per person. The same holds true for the age composition, which is relatively similar across countries as well (see Online Appendix Figure B.1). The educational and sectoral composition in turn vary significantly across countries. The left panel of Figure 7 shows that the US stands out with the lowest share of low educated individuals with less than 10 percent, compared to nearly 40 percent in Southern Europe. The reversed pattern prevails for high educated individuals, while the fraction of young adults still enrolled in education is rather similar across regions.

The right panel of Figure 7 shows the sectoral composition across regions. The share of em-

Figure 7: Demographic Structure

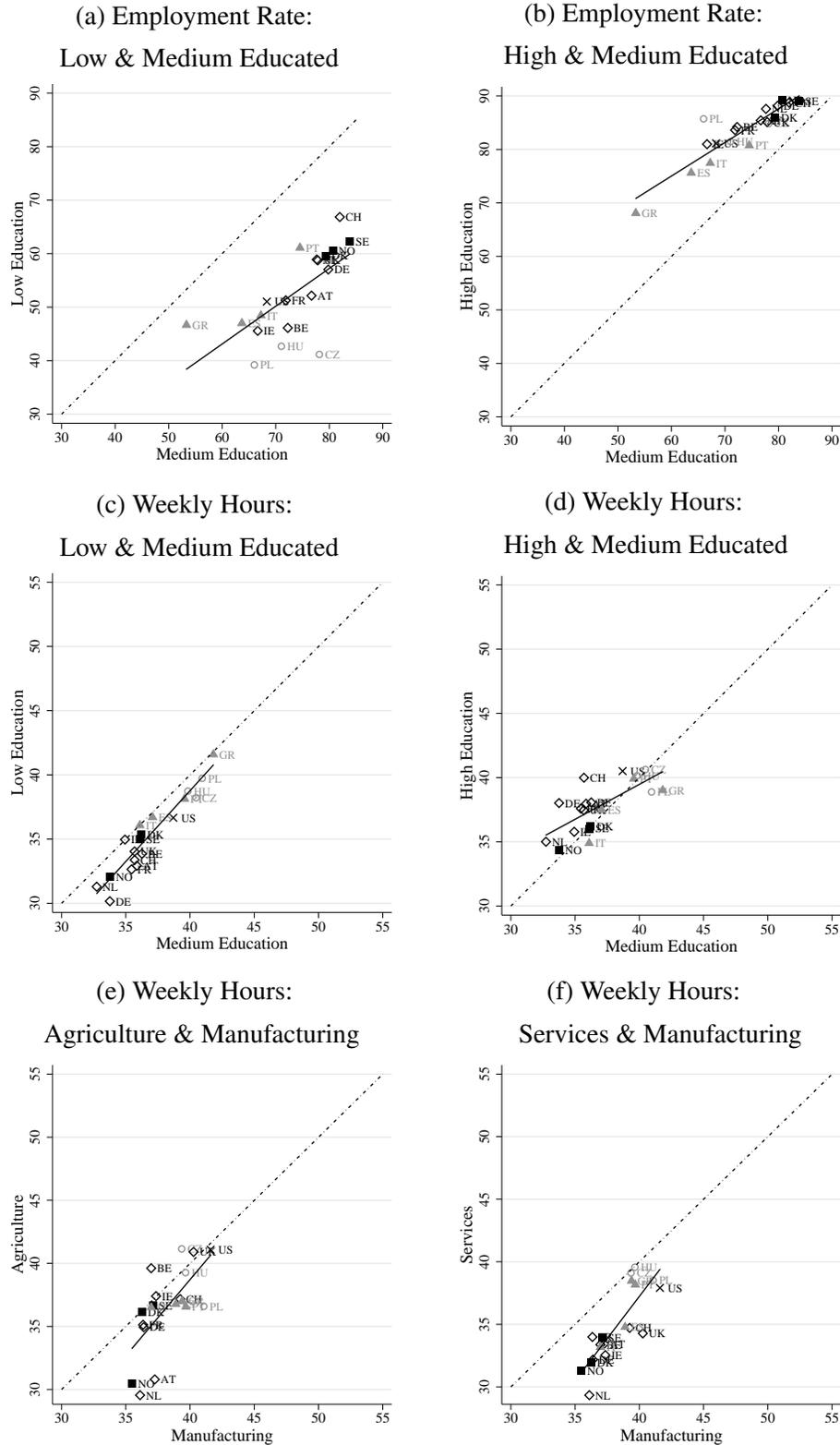


Note: Each bar shows for the years 2013 to 2015 (a) the fraction (in %) of the 15 to 64 year old population in each education category; (b) the fraction (in %) of the 15 to 64 year old employed population in each sectoral category.

employed people working in services is highest in the US with about 80 percent and lowest in Eastern Europe with around 60 percent. The low share in services in Eastern and Southern Europe is offset by both a higher share of individuals working in manufacturing, and a slightly higher share working in agriculture, with the latter being quite low everywhere.

While the sectoral and the educational compositions thus differ significantly across countries, this only has an effect in the disaggregate decomposition if different educational groups show different labor market behavior within a country, and individuals in different sectors work different hours. Panels (a) to (d) in Figure 8 demonstrate why the large differences in the educational composition accounts for a large fraction of the Europe-US hours gap. The differences in weekly hours worked (panels (c) and (d)) by low, medium, and high educated individuals within a country are rather small, but the differences in employment rates are substantial. On average across all countries, medium educated individuals have a 21 percentage points higher employment rate than low educated individuals, but a 10 percentage points lower one than high educated individuals. Thus, the large differences in educational shares between Europe and the US, and in employment rates by education within countries account for around one third to one half of the Europe-US hours gap.

Figure 8: Labor Supply Broken Down by Education or Sector



Panels (a) and (b) plot the country-specific employment rates for different levels of education against each other. Panels (c) and (d) do the same for weekly hours. Panels (e) and (f) compare average weekly hours for different sectors. Each average is calculated for the population aged 15 to 64 for 2013 to 2015. The dash-dotted line is the 45 degree line.

By contrast, panels (e) and (f) of Figure 8 show that the differences in hours worked across sectors are quite small and therefore do not contribute meaningfully to the aggregate Europe-US hours gap. Individuals working in the manufacturing sector work on average 3.4 more weekly hours than individuals in the service sector, and the differences between manufacturing and agriculture are even smaller. Combined with a larger service sector in the US, these sectoral hours differences predict minimally higher hours in Europe than in the US. As a reminder, in our decomposition exercise the sectoral composition can only affect hours worked per person through differences in weekly hours worked *per employed* between sectors, since we cannot assign non-employed individuals to a sector. Rogerson (2008) stresses the importance of an underdeveloped service sector in Europe for accounting for low hours worked *per person* there. In his model, high tax rates in Europe induce individuals to specialize in home production, leading to a stagnating service sector instead of a growing one as in the US, since services produced in the market are closer substitutes to services produced at home than goods produced in the market to goods produced at home. As far as this channel works through drawing less people into employment, it is not captured in our decomposition, which by construction can only handle sectoral differences in weekly hours worked.

### ***Residual Differences in Employment Rates and Weekly Hours Worked***

After investigating the role of the demographic composition for the Europe-US hours difference, we analyze in this section how the differences in employment rates and weekly hours worked not attributed to differences in the demographic composition, i.e. the residual differences  $e$  and  $h$  in Figure 6, vary by demographic subgroups. Given the negligible role of sectoral differences in hours worked per employed, we focus our analysis solely on gender, age, and education, and present results again for the four European regions.<sup>23</sup> Specifically, we construct for each country

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<sup>23</sup>In contrast to the decomposition exercise in the previous section, we include the countries with the smallest differences in each country group. The exception is Portugal in Southern Europe. Similarly, we exclude the group “enrolled in education” for the young. We show the results for both cases in Online Appendix Section B.4.

employment rates and weekly hours worked for all cells combining age, education, and gender (i.e. 18 cells), pool all countries for each region, and regress the employment rate differences (in percentage points) and weekly hours worked differences (in percent) relative to the US on a set of dummy variables for age, education, and gender.<sup>24</sup> The baseline category are medium educated men of core ages (25 to 54), which typically show a large attachment to the labor market and are one of the largest groups in each country. The corresponding coefficient is shown as the first bar in each panel of Figure 9. To give a concrete example, consider panels (a) and (b) for Scandinavia. Scandinavian medium educated men of core ages have an 8 percentage point higher employment rate than their US counterpart, but work around 10% fewer hours per week. The remaining bars give the sum of the constant (the baseline category) and the coefficient on the respective dummy, and represent the difference to the US for this group. E.g. the second bar shows that Scandinavian medium educated women of core ages have a 14 percentage point higher employment rate than their US counterpart, and 13 percent lower weekly hours worked. As panels (a) and (b) show, both higher employment rates and lower weekly hours worked in Scandinavia than in the US are driven by all groups.

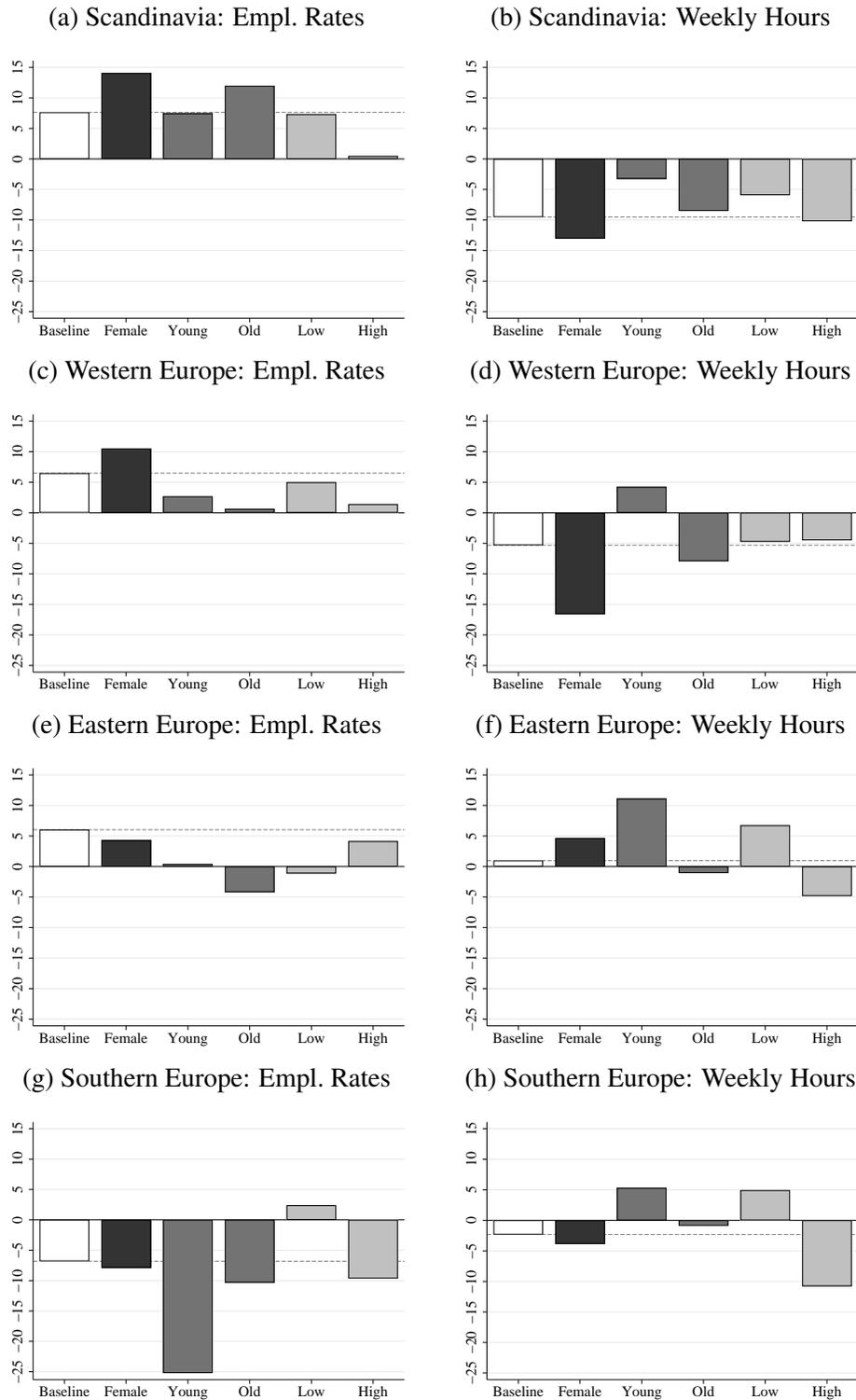
Western Europe is similar to Scandinavia. Controlling for the demographic composition, employment rates predict higher hours worked per person than in the US, and weekly hours worked lower ones, and this is quite homogeneously true for different demographic subgroups, see panels (c) and (d). As in Scandinavia, women stand somewhat out with the largest differences to the US, especially in weekly hours, and younger individuals with the smallest differences to their US counterpart in weekly hours (which even turns positive in Western Europe).

In contrast to Scandinavia and Western Europe, Eastern European employment rates account for a non-negligible part of the lower hours worked per person than in the US. This is predominantly driven by older individuals, as well as low-educated ones, see panel (e). The on average

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<sup>24</sup>Table B.6 in the Online Appendix reports for each group and region the corresponding differences. Table B.7 shows the actual regression coefficients and thus allows to construct differences to the US for groups that deviate from the baseline group by more than one category (e.g. high-educated older women).

Figure 9: Differences between Europe and the US by Demographic Group



Note: The graphs show results from a regression of employment rate differences (in percentage points, left panels) and weekly hours worked differences (in percent, right panels) between different European regions and the US, broken down by gender, age, and education, on a female dummy, dummies for young (15-24 years) and old (55-64 years), and dummies for low and high education. The regressions are run separately by region, and Portugal and the educational category “still enrolled” are excluded. The white bars and gray dashed lines show the estimated constant, representing the average difference for men aged 25-54 with medium education. The other graphs reflect the sum of the constant and the coefficients for the different dummies. The exact regression results can be found in Table B.7. All results are for the years 2013 to 2015.

higher weekly hours worked in Eastern Europe compared to the US originate mostly from women, the young, and the low educated. Among all country groups, the residual differences in Eastern Europe are the most heterogeneous across different demographic groups.

Southern Europe (excluding Portugal) is the only region in which both employment rates and weekly hours contribute positively to the lower hours worked per person than in the US. As panels (g) and (h) show, this is again quite homogeneously true for all groups, except for the low educated. Low educated core aged men exhibit higher employment rates and higher weekly hours than their US counterparts. However, remember that the share of low educated individuals is exceptionally high in Southern Europe. Thus, there is likely a more positive selection into this group than in the US. At the same time, the young stand out with exceptionally low employment rates compared to the US, a result of the high youth unemployment rates in Southern European countries.

Summing up the patterns across groups and regions, we find that for Scandinavia, Western Europe, and Southern Europe the basic patterns are quite homogeneous across groups: in Scandinavia and Western Europe, all groups tend to exhibit higher employment rates and lower weekly hours than their US counterparts. However, being female increases the differences to the US in both margins. In Southern Europe, all demographic groups tend to exhibit both lower employment rates and lower weekly hours than their US counterparts, with the exception of the low educated. Eastern Europe features more heterogeneities across the groups than the other regions. This might partly be due to a transition effect after the abolishment of socialism. E.g. older individuals stand out with low employment rates and low hours, which could be due to a skill mismatch after the transition.

## **V. Conclusion**

This paper constructs “new data” on hours worked per person for the US and 18 European countries from national labor force surveys for the years 1983 through 2015. While using labor force survey data for cross-country studies of employment and hours worked is not “new”, our contribution

is to apply the harmonization procedure suggested by Pilat (2003), and to systematically analyze its importance. This harmonization procedure addresses the two main challenges for measuring hours worked consistently across countries and over time from labor force surveys: the different sampling of reference weeks over the year, and the underreporting of vacation weeks. We show that we overcome both problems by obtaining hours worked in a *non-vacation* week, i.e. a work week without a reduction in work time because of a public holiday or annual leave, and requiring consistency of annual leave and public holidays in the micro data with the country-wide average.

The richness of our micro data allows us to move beyond the aggregate hours worked differences between Europe and the US and to provide “new answers” regarding the origins of these differences. We provide a decomposition of the Europe-US hours gap into the roles played by weeks worked, employment rates, weekly hours worked per employed in a non-vacation week, and the demographic and sectoral compositions. Weeks worked and the demographic composition, driven by the educational composition, both account for between one quarter and one half of the lower hours per person in Europe than in the US. In addition, lower weekly hours worked are important drivers of the lower hours in Scandinavia and Western Europe, where employment rates alone would imply higher hours per person than in the US. These differences are present for all demographic groups considered in our analysis, with women showing the largest differences. For Eastern Europe, employment rate differences account for roughly one quarter of the hours per person gap with the US, while weekly hours predict higher hours than in the US. Only in Southern Europe (excluding Portugal), employment rate and weekly hours differences both contribute positively to the hours gap to the US, with employment rates being much more important. All demographic groups tend to exhibit both lower employment rates and lower weekly hours in Southern Europe than their US counterparts, with the exception of the low educated.

Our results indicate new avenues for the research agenda of the study of hours worked differences across Europe and the US. First, they show that the higher number of vacation weeks in Europe is a significant driver of lower hours, and that longer vacation weeks are hardly associated

with either higher employment rates or longer weekly hours worked. This might indicate that the number of work weeks is determined by other driving factors than those determining employment rates and weekly hours worked, and raises the question whether they should be modeled as a choice variable for the individual. One could instead follow e.g. a specification as in [Kaplan \(2012\)](#), which takes the number of work weeks as exogenously given. Second, the educational composition matters significantly in accounting for cross-country differences in hours worked through its effect on employment rates. Understanding why different education groups exhibit markedly different employment rates but similar weekly hours worked, and why different countries feature very different educational compositions, will thus be a helpful step in understanding the origins of hours worked differences across countries. For example, [Restuccia and Vandenbroucke \(2014\)](#) analyze the role of longevity and economic growth for educational attainment and hours worked. [Guvenen et al. \(2014\)](#) show that progressive taxation distorts the incentives to invest in human capital, and differences in progressivity across countries can explain a sizable part of the differences in the Europe-US hours per person gap for men. However, since these papers do not distinguish between both margins, their theories remain silent about the very different education gradients between employment and hours worked per employed. Other candidate explanations affecting the incentives to work differentially by education are welfare systems, or different degrees of regulations in the labor markets for low and high educated, e.g. due to a minimum wage being binding for the poor. Yet, one would have to investigate whether these factors influence the extensive margin more than the intensive one.

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